

# Lecture 2b: Survey weights

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Advanced Methods of Social Research (SOCL 420)

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Source: Treiman, Donald J. 2009. Quantitative Data Analysis: Doing Social Research to Test Ideas. San Francisco: Jossey-Bass. Chapter 9 (pp. 195–224).



# Outline

- Inferential statistics
- Survey weights
- Weight options in Stata
- Complex sample cluster design
- Weights in the General Social Survey (GSS)
- Examples



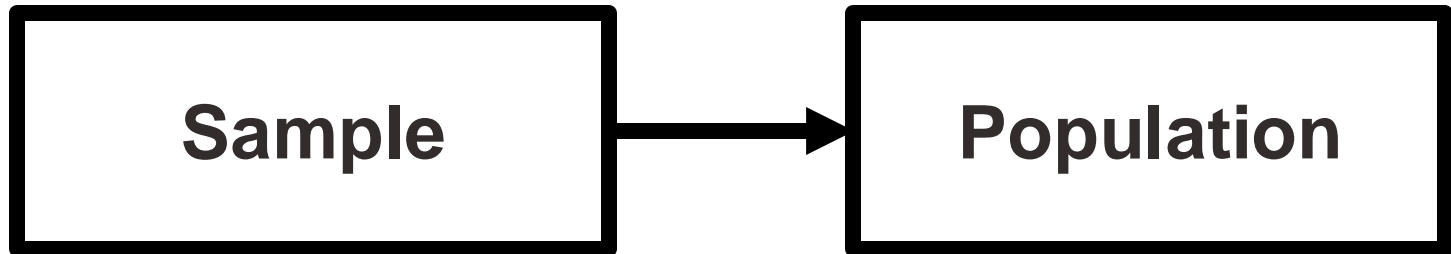
# Inferential statistics

- Social scientists need inferential statistics
  - They almost never have the resources or time to collect data from every case in a population
- Inferential statistics uses data from samples to make generalizations about populations
  - **Population** is the total collection of all cases in which the researcher is interested
  - **Samples** are carefully chosen subsets of the population
- With proper techniques, generalizations based on samples can represent populations

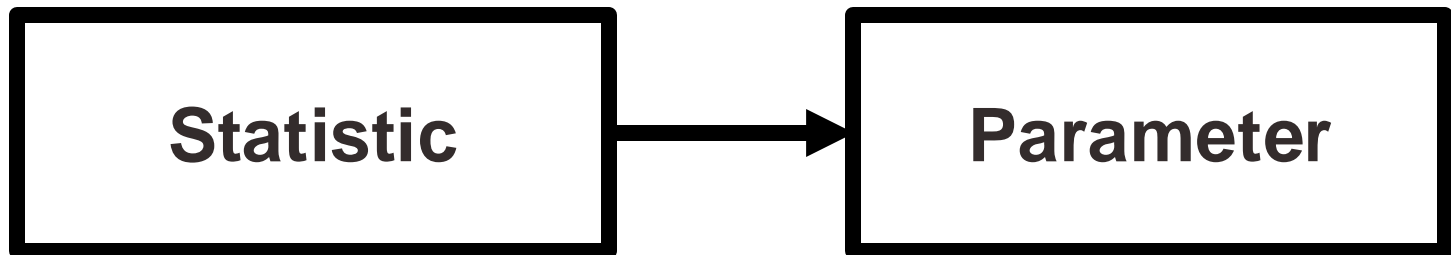


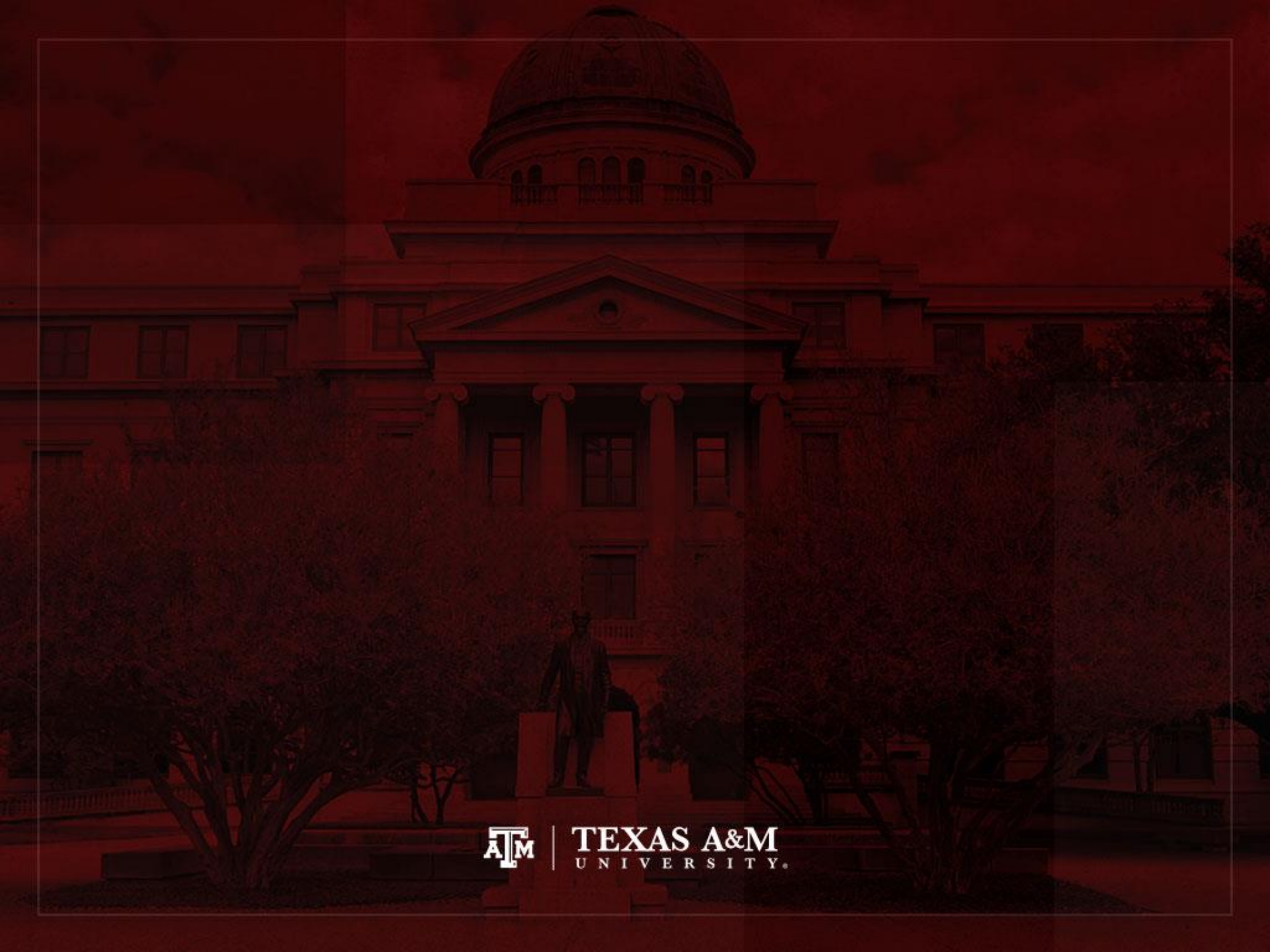
# Basic logic and terminology

- Information from samples is used to estimate information about the population



- Statistics: characteristics of samples
- Parameters: characteristics of populations
- Statistics are used to estimate parameters





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# Survey weights

Name	Number of observations collected in the survey	Weight to expand to population size	Weight to maintain sample size
José	1	4	0.8
Maria	1	6	1.2
Total	2	10	2

**Survey weight =**

**Population weight \* (Sum of survey weights / Sum of population weights)**



# Weights for tables

- When we use a sample to estimate the absolute number of people
  - For an area
  - For a specific sub-group
  - We use weights to expand to population size
- If we use a sample to estimate the proportion of people in a specific sub-group
  - And we are not concerned with the absolute value
  - We use weights to maintain the sample size (we focus on percentages)



# Weights for regressions

- In a simple linear regression, the test of statistical significance for a  $\beta$  coefficient ( $t$ -test) is estimated as

$$t = \frac{\hat{\beta}}{SE_{\hat{\beta}}} = \frac{\hat{\beta}}{\sqrt{\frac{MSE}{S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{RSS}{df * S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{\sum_i (y_i - \hat{y}_i)^2}{(n - 2) \sum_i (x_i - \bar{x})^2}}}$$

- $SE_{\hat{\beta}}$ : standard error of  $\beta$
- $MSE$ : mean squared error =  $RSS / df$
- $RSS$ : residual sum of squares =  $\sum_i (y_i - \hat{y}_i)^2 = \sum_i \hat{e}_i^2$
- $df$ : degrees of freedom =  $n-2$  for simple linear regression
  - 2 statistics (slope and intercept) are estimated to calculate sum of squares
- $S_{xx}$ : corrected sum of squares for  $x$  (total sum of squares)





# Weights for regressions

- If we use a weight that expands to the population size ( $N$ ) on regressions
  - We would be incorrectly informing the statistical software that we have a sample with enormous size
  - This would artificially increase the test of statistical significance for the coefficient

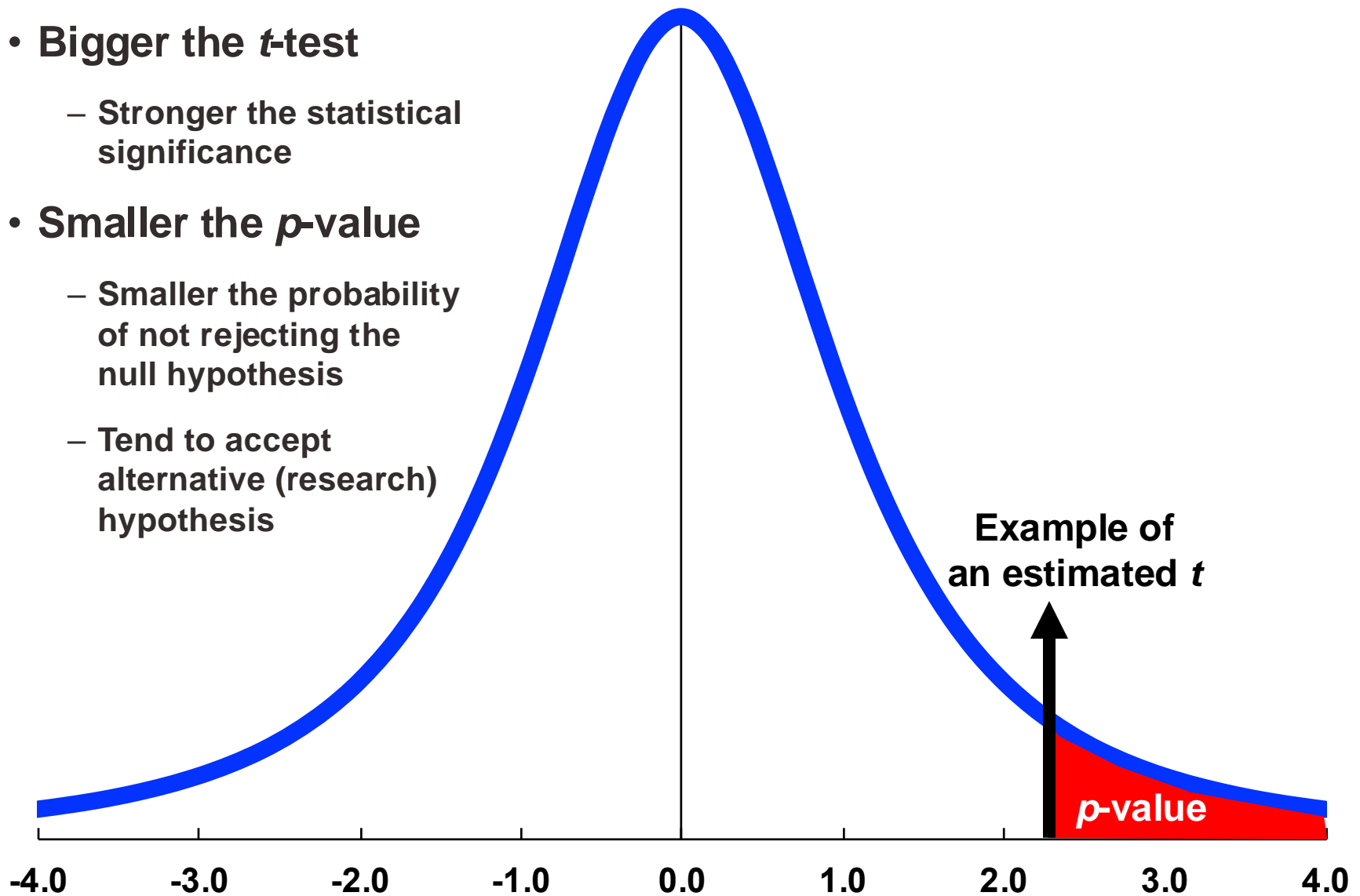
$$\uparrow t = \frac{\hat{\beta}}{SE_{\hat{\beta}}} = \frac{\hat{\beta}}{\sqrt{\frac{MSE}{S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{MSE}{S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{\sum_i (y_i - \hat{y}_i)^2}{(n-2) \sum_i (x_i - \bar{x})^2}}}$$

The equation shows the derivation of the t-statistic. A red arrow points up to the 't' and another red arrow points down to the denominator of the final fraction.

- We have to inform the weight related to the sample design, but we should maintain the sample size ( $n$ )

# $t$ distribution ( $df = 2$ )

- Bigger the  $t$ -test
  - Stronger the statistical significance
- Smaller the  $p$ -value
  - Smaller the probability of not rejecting the null hypothesis
  - Tend to accept alternative (research) hypothesis



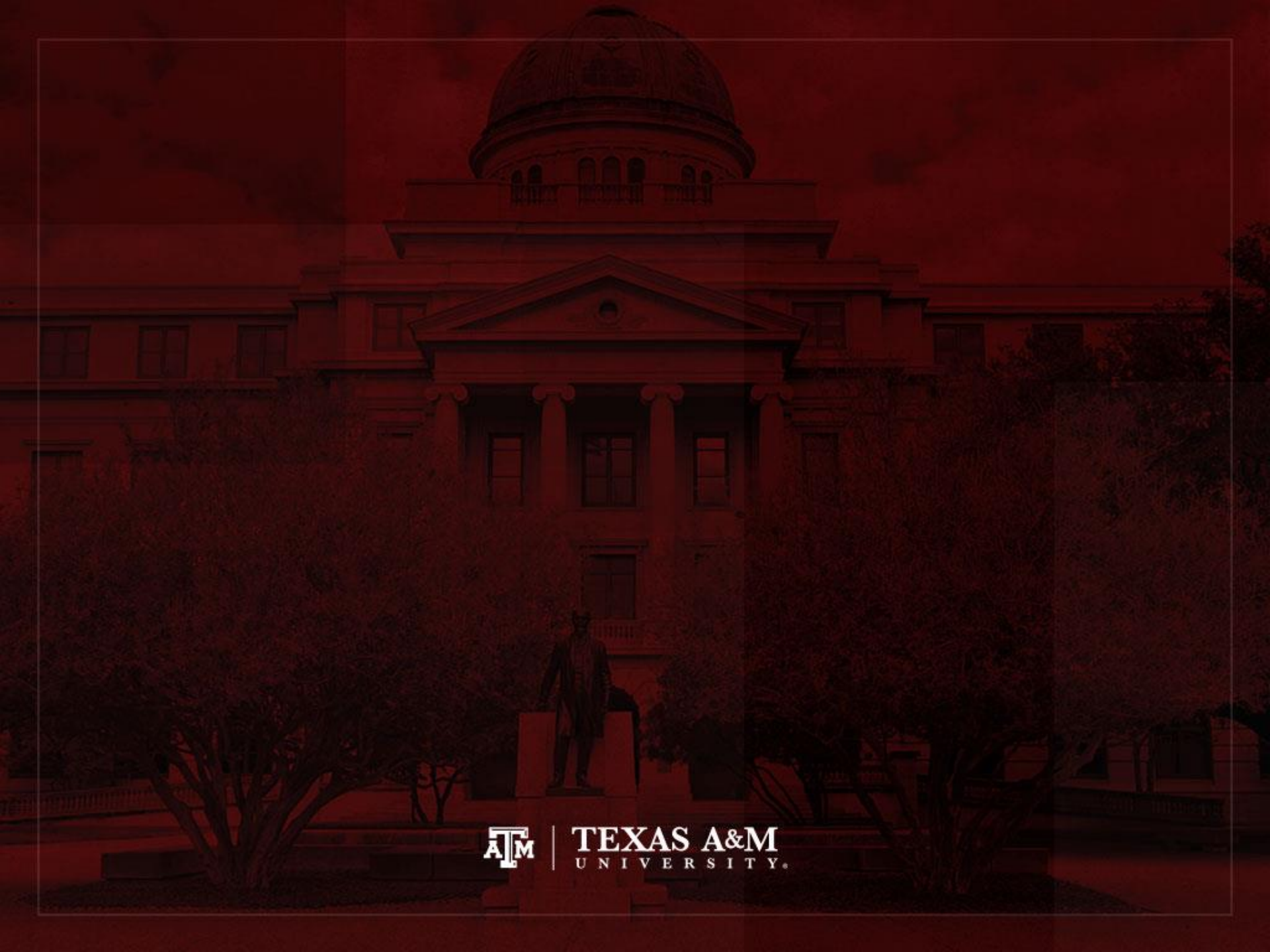
# Decisions about hypotheses

Hypotheses	$p < \alpha$	$p > \alpha$
Null hypothesis ( $H_0$ )	Reject	Do not reject
Alternative hypothesis ( $H_1$ )	Accept	Do not accept

- **$p$ -value** is the probability of not rejecting the null hypothesis
- If a statistical software gives only the two-tailed  $p$ -value, divide it by 2 to obtain the one-tailed  $p$ -value

Significance level ( $\alpha$ )	Confidence level (success rate)
0.10 (10%)	90%
0.05 (5%)	95%
0.01 (1%)	99%
0.001 (0.1%)	99.9%





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# Weight options in Stata

- Frequency weight (fweight)
- "Importance" weight (iweight)
- Analytic weight (aweight)
- Sampling weight (pweight)



# Extract of 2018 ACS microdata

	year	strata	cluster	perwt	hhwt	sex	age	income
1	2018	360248	2.018012e+12	56.00	56.00	Male	46	28000
2	2018	360248	2.018012e+12	51.00	51.00	Male	20	5000
3	2018	360248	2.018012e+12	76.00	76.00	Female	84	0
4	2018	360248	2.018012e+12	55.00	55.00	Female	18	1200
5	2018	360248	2.018012e+12	143.00	143.00	Female	56	1500
6	2018	360248	2.018012e+12	198.00	198.00	Male	31	10000
7	2018	360248	2.018012e+12	48.00	48.00	Female	19	2000
8	2018	360248	2.018012e+12	48.00	48.00	Male	25	7000
9	2018	360248	2.018012e+12	65.00	65.00	Female	18	0
10	2018	360248	2.018012e+12	53.00	53.00	Female	18	15000
11	2018	360248	2.018012e+12	17.00	17.00	Male	63	0
12	2018	360248	2.018012e+12	39.00	39.00	Female	18	4000
13	2018	360248	2.018012e+12	104.00	104.00	Male	21	1000
14	2018	360248	2.018012e+12	200.00	200.00	Male	40	80000
15	2018	360248	2.018012e+12	20.00	20.00	Male	33	0
16	2018	360248	2.018012e+12	59.00	59.00	Male	19	2900
17	2018	360248	2.018012e+12	56.00	56.00	Male	55	0
18	2018	360248	2.018012e+12	77.00	77.00	Male	18	9000
19	2018	360248	2.018012e+12	16.00	16.00	Female	41	1100
20	2018	360248	2.018012e+12	46.00	46.00	Male	33	0

Source: 2018 American Community Survey (ACS).

# Extract of 2022 GSS microdata

	year	vstrat	vpsu	wtssnrps	sex	age	conrinc
1	2022	2661	1	.3009902626114151492	female	72	63300
2	2022	2661	1	.741964891046645203	male	80	iap
3	2022	2661	1	1.363834172908697351	female	57	28485
4	2022	2661	1	1.202291262405986272	female	23	3481.5
5	2022	2661	1	1.452537303169801897	male	62	iap
6	2022	2661	1	1.357217051758134341	male	27	34815
7	2022	2661	1	.4122643975892986146	female	20	23737.5
8	2022	2661	1	.6095341671847818166	male	47	52222.5
9	2022	2661	1	.4756119451858641378	female	31	10286.25
10	2022	2661	1	.4785392835774277542	female	72	iap
11	2022	2661	1	1.122700132546621044	female	57	iap
12	2022	2661	1	.3201873333462621485	female	25	20572.5
13	2022	2661	1	1.100824120020779118	female	35	28485
14	2022	2661	2	.3851760620571728078	female	36	316.5
15	2022	2661	2	.7969904442288312518	female	41	iap
16	2022	2661	2	.2670620034929151454	male	65	iap
17	2022	2661	2	.6600438649618246556	female	20	iap
18	2022	2661	2	1.044982188718633909	female	80	20572.5
19	2022	2661	2	.2466453269085958078	female	35	7121.25
20	2022	2661	2	.3875467776907561834	female	89 or older	iap
21	2022	2661	2	1.811167810753548713	female	34	2215.5
22	2022	2661	2	1.358934837899157033	female	55	iap
23	2022	2661	2	.5032689647947735656	male	65	20572.5
24	2022	2661	2	1.101277115765816328	female	60	28485
25	2022	2661	2	1.357217051758134341	male	28	316.5

# Frequency weight

- **FWEIGHT**

- Expands survey size to the population size
- Indicates the number of duplicated observations
- Used on tables to generate frequencies
- Can be used in frequency distributions only when weight variable is discrete (no fractional numbers)

```
tab x [fweight = weight]
```





# "Importance" weight

- **IWEIGHT**

- Indicates the "importance" of the observation in some vague sense
- Has no formal statistical definition
- Any command that supports iweights will define exactly how they are treated
- Intended for use by programmers who want to produce a certain computation
- Can be used in frequency distributions even when weight variable is continuous (fractional numbers)

```
tab x [iweight = weight]
```



# Analytic weight

- **AWEIGHT**

- Inversely proportional to the variance of an observation
- Variance of the  $j$ th observation is assumed to be  $\sigma^2/w_j$ , where  $w_j$  are the weights
- For most Stata commands, the recorded scale of aweights is irrelevant
- Stata internally rescales frequencies, so sum of weights equals sample size

```
tab x [aweight = weight]
```

```
regress y x1 x2 [aweight = weight]
```



# More about analytic weight

- Observations represent averages and weights are the number of elements that gave rise to the average

group	x	y	n
1	3.5	26.0	2
2	5.0	20.0	3

- Instead of

group	x	y
1	3	22
1	4	30
2	8	25
2	2	19
2	5	16

- Usually, survey data is collected from individuals and households (not as averages)
  - Thus, aweights are not appropriate for most cases



# Sampling weight

- **PWEIGHT**

- Denote the inverse of the probability that the observation is included due to the sampling design
- Variances, standard errors, and confidence intervals are estimated with a more precise procedure
- Indicated for statistical regressions to estimate robust standard errors
  - Obtain unbiased standard errors of OLS coefficients under heteroscedasticity (i.e., residuals not randomly distributed)
  - Robust standard errors are usually larger than conventional ones

`regress y x1 x2 [pweight = weight]`



# Summary of Stata weights

## WEIGHTS IN FREQUENCY DISTRIBUTIONS

Weight unit of measurement	Expand to population size	Maintain sample size
Discrete	fweight	aweight
Continuous	iweight	

## WEIGHTS IN STATISTICAL REGRESSIONS should maintain sample size

Robust standard error	Adjusted R <sup>2</sup> , TSS, ESS, RSS
pweight svy: reg y x	aweight
reg y x, vce(robust) reg y x, vce(cluster area)	outreg2



# Example of 2018 ACS weight

```
. sum perwt, d
```

Person weight

---

	Percentiles	Smallest		
1%	<b>10</b>	<b>1</b>		
5%	<b>19</b>	<b>1</b>		
10%	<b>29</b>	<b>1</b>	Obs	<b>3,214,539</b>
25%	<b>52</b>	<b>1</b>	Sum of wgt.	<b>3,214,539</b>
50%	<b>80</b>		Mean	<b>101.7774</b>
		Largest	Std. dev.	<b>83.93534</b>
75%	<b>124</b>	<b>1916</b>		
90%	<b>195</b>	<b>1990</b>	Variance	<b>7045.14</b>
95%	<b>263</b>	<b>2097</b>	Skewness	<b>2.845116</b>
99%	<b>427</b>	<b>2313</b>	Kurtosis	<b>17.99265</b>

# Example of 2018 ACS weight

. tab sex

Sex	Freq.	Percent	Cum.
Male	1,574,618	48.98	48.98
Female	1,639,921	51.02	100.00
Total	3,214,539	100.00	

. tab sex [fweight=perwt]

Sex	Freq.	Percent	Cum.
Male	161,072,404	49.23	49.23
Female	166,095,035	50.77	100.00
Total	327,167,439	100.00	

. tab sex [iweight=perwt]

Sex	Freq.	Percent	Cum.
Male	161,072,404	49.23	49.23
Female	166,095,035	50.77	100.00
Total	327,167,439	100.00	

. tab sex [aweight=perwt]

Sex	Freq.	Percent	Cum.
Male	1,582,595	49.23	49.23
Female	1,631,944	50.77	100.00
Total	3,214,539	100.00	



# Example of 2021 GSS weight

. sum wtssnrps, d

person post-stratification weight, nonrespondents  
adjusted

---

	Percentiles	Smallest		
1%	.243687	.1723802		
5%	.30024	.1738938		
10%	.4057674	.1926333	Obs	4,032
25%	.5423563	.2104285	Sum of wgt.	4,032
50%	.8183308		Mean	1
		Largest	Std. dev.	.7260472
75%	1.212269	6.51434		
90%	1.798724	6.903664	Variance	.5271445
95%	2.27083	7.218392	Skewness	2.825826
99%	3.986099	7.557038	Kurtosis	15.89999



# Example of 2021 GSS weight

```
. tab sex, m
```

respondents sex	Freq.	Percent	Cum.
male	1,736	43.06	43.06
female	2,204	54.66	97.72
.i	19	0.47	98.19
.n	71	1.76	99.95
.s	2	0.05	100.00
Total	4,032	100.00	

```
. tab sex [fweight=wtssnrps], m
```

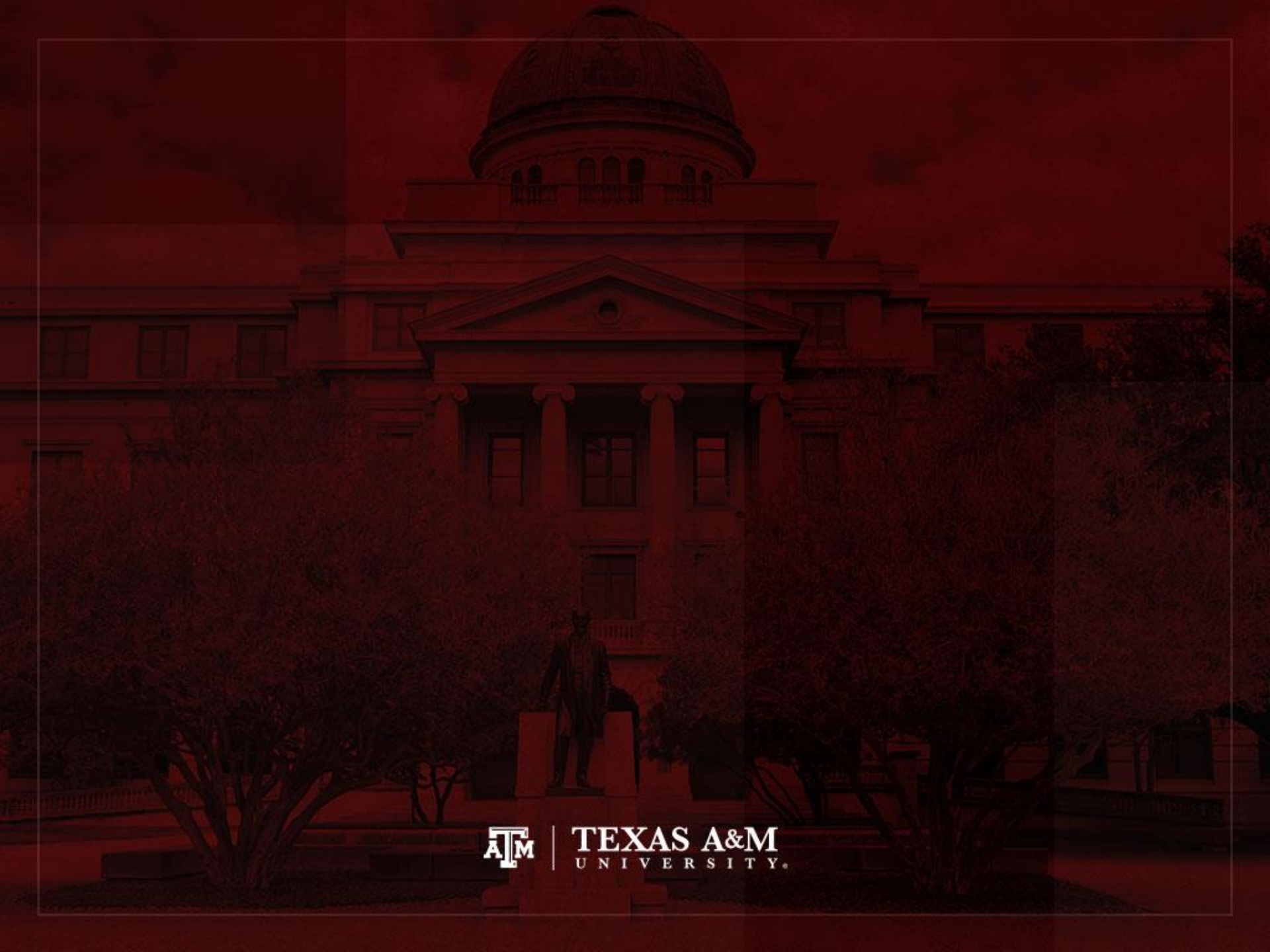
may not use noninteger frequency weights  
r(401);

```
. tab sex [iweight=wtssnrps], m
```

respondents sex	Freq.	Percent	Cum.
male	1,904.2566	47.23	47.23
female	1,993.21543	49.43	96.66
.i	18.1122752	0.45	97.11
.n	113.299832	2.81	99.92
.s	3.11586052	0.08	100.00
Total	4,032	100.00	

```
. tab sex [aweight=wtssnrps], m
```

respondents sex	Freq.	Percent	Cum.
male	1,904.2566	47.23	47.23
female	1,993.21543	49.43	96.66
.i	18.1122752	0.45	97.11
.n	113.299832	2.81	99.92
.s	3.11586052	0.08	100.00
Total	4,032	100.00	



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# Complex sample cluster design

- To calculate standard errors correctly, variables for sample cluster design must be used
  - Without design variables, Stata will assume a simple random sample and underestimate standard errors
- Strata are created based on the lowest level of geography available in each sample
  - We use additional statistical techniques that account for the complex sample design to produce correct standard errors and statistical tests



# Cluster design for tables

- If we want to estimate a confidence interval for a sample statistic (mean or proportion), we need to inform the complex survey design
- **Confidence interval** is a range of values used to estimate the true population parameter
- **Confidence level** is the success rate of the procedure to estimate the confidence interval
- Larger confidence levels generate larger confidence intervals



# Confidence level, $\alpha$ , and $Z$

Confidence level (1 - $\alpha$ ) * 100	Significance level alpha ( $\alpha$ )	$\alpha / 2$	Z score
90%	0.10	0.05	$\pm 1.65$
<b>95%</b>	<b>0.05</b>	<b>0.025</b>	<b><math>\pm 1.96</math></b>
99%	0.01	0.005	$\pm 2.58$
99.9%	0.001	0.0005	$\pm 3.32$
99.99%	0.0001	0.00005	$\pm 3.90$



# Confidence intervals from samples

*c.i. = sample estimate  $\pm$  margin of error*

*c.i. = sample estimate  $\pm$  score of confidence level \* standard error*

- Sample mean ( $\bar{x}$ ), standard deviation (s),  $n < 30$

$$c.i. = \bar{x} \pm t \left( \frac{s}{\sqrt{n}} \right) \quad df = n - 1$$

- Sample mean ( $\bar{x}$ ), standard deviation (s),  $n \geq 30$

$$c.i. = \bar{x} \pm Z \left( \frac{s}{\sqrt{n - 1}} \right)$$

- Sam. proportion ( $P_s$ ), pop. proportion ( $P_u$ ),  $n \geq 30$

$$c.i. = P_s \pm Z \sqrt{\frac{P_u(1 - P_u)}{n}}$$



# Cluster design for regressions

- We also need to inform cluster design for regressions, because the  $t$ -test utilizes standard errors

$$t = \frac{\hat{\beta}}{SE_{\hat{\beta}}} = \frac{\hat{\beta}}{\sqrt{\frac{MSE}{S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{RSS}{df * S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{\sum_i (y_i - \hat{y}_i)^2}{(n - 2) \sum_i (x_i - \bar{x})^2}}}$$

- $SE_{\hat{\beta}}$ : standard error of  $\hat{\beta}$
- $MSE$ : mean squared error =  $RSS / df$
- $RSS$ : residual sum of squares =  $\sum_i (y_i - \hat{y}_i)^2 = \sum_i \hat{e}_i^2$
- $df$ : degrees of freedom =  $n-2$  for simple linear regression
- $S_{xx}$ : corrected sum of squares for  $x$  (total sum of squares)



# Cluster design & standard error

- Sample cluster designs underestimate standard errors, because they tend to select individuals with more similar characteristics from the same clusters
  - Simple random samples would provide more variation (higher standard errors), because they give the same chance of selection for all individuals in the population
- When we inform the cluster design, the standard error tends to increase and statistical significance decreases

$$\downarrow t = \frac{\hat{\beta}}{\uparrow SE_{\hat{\beta}}} = \frac{\hat{\beta}}{\sqrt{\frac{MSE}{S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{RSS}{df * S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{\sum_i (y_i - \hat{y}_i)^2}{(n - 2) \sum_i (x_i - \bar{x})^2}}}$$







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# Weights in GSS

- The General Social Survey (GSS) targets the adult population (18+) living in U.S. households
- Due to the adoption of the sub-sampling design of non-respondents, a weight must be employed when using the GSS 2004 and after
- These are the weight variables
  - WTSS, WTSSPS
  - WTSSNR, WTSSNRPS
  - WTSSALL
- They all maintain the original sample size, even in frequency distributions with “iweight”



# WTSS, WTSSPS

- WTSS variable takes into consideration
  - Sub-sampling of non-respondents
  - Number of adults in the household
  - Starting in 2021, this weight is called WTSSPS, which refers to probability sampling
- In years prior to 2004, a value of one is assigned to all cases, so they are effectively unweighted
  - Number of adults can be utilized to make this adjustment for years prior to 2004



# WTSSNR, WTSSNRPS

- WTSSNR variable takes into consideration
  - Sub-sampling of non-respondents
  - Number of adults in the household
  - Differential non-response across areas
  - Starting in 2021, this weight is called WTSSNRPS, which refers to probability sampling with non-response adjustment
- In years prior to 2004, a value of one is assigned to all cases, so they are effectively unweighted
  - Number of adults can be utilized to make this adjustment for years prior to 2004
  - Area non-response adjustment is not possible



# WTSSALL

- WTSSALL takes WTSS and applies an adult weight to years before 2004
- The weight value of WTSSALL is the same as WTSS for 2004 and after
- Starting in 2021, WTSSALL is not provided due to changes in data collection methods during the COVID-19 pandemic

# Multi-year analysis

- For multi-year analysis including 2021 data and beyond, use the following weights...
- Pre-2021 data
  - Use WTSSALL, which accounts for the probability of selection, subsampling, and the number of adults in the household
- 2021 and beyond
  - Use WTSSNRPS, a post-stratification weight that adjusts for survey design and nonresponse, aligning the sample with U.S. Census Bureau estimates



# Create a unified weight variable

- In Stata, you can create a unified weight variable for multi-year analysis

```
*Generate a new weight variable  
gen weightfinal=.
```

```
*Pre-2021: Use WTSSALL
```

```
replace weightfinal=wtssall if year>=1972 & year<=2018
```

```
*2021 and beyond: Use WTSSNRPS
```

```
replace weightfinal=wtssnrps if year>=2021
```

```
*Descriptive statistics with the new weight variable
```

```
tab x [aweight = weightfinal]
```

```
sum x [aweight = weightfinal]
```



# GSS has a cluster sample

([https://gssdataexplorer.norc.org/gss\\_stdError](https://gssdataexplorer.norc.org/gss_stdError))

- First- and second-stage units are selected with probabilities proportional to size
  - Size is defined by number of housing units
- Third-stage units (housing units) are selected to be an equal-probability sample
  - This results in roughly the same number of housing units selected per second-stage sampling unit





# GSS variables for cluster design

([https://gssdataexplorer.norc.org/gss\\_stdError](https://gssdataexplorer.norc.org/gss_stdError))

- There are two design variables
  - VSTRAT
  - VPSU
- First-stage unit
  - VSTRAT: Variance Stratum
  - National Frame Areas (NFAs): one or more counties
- Second-stage unit
  - VPSU: Variance Primary Sampling Unit
  - Segments: block, group of blocks, or census tract



# GSS complex sample design

([https://gssdataexplorer.norc.org/gss\\_stdError](https://gssdataexplorer.norc.org/gss_stdError))

- Code to account for GSS sample design in Stata  
`svyset [weight=weightfinal], strata(vstrat)  
psu(vpsu) singleunit(scaled)`
  - “weightfinal” was previously created using WTSSALL (pre-2021) and WTSSNRPS (2021 and beyond)
- After “svyset,” you should indicate the survey design with the option “svy” for commands that estimate standard errors

`svy: mean y`

`svy: reg y x1 x2`



# Strata with single sampling unit

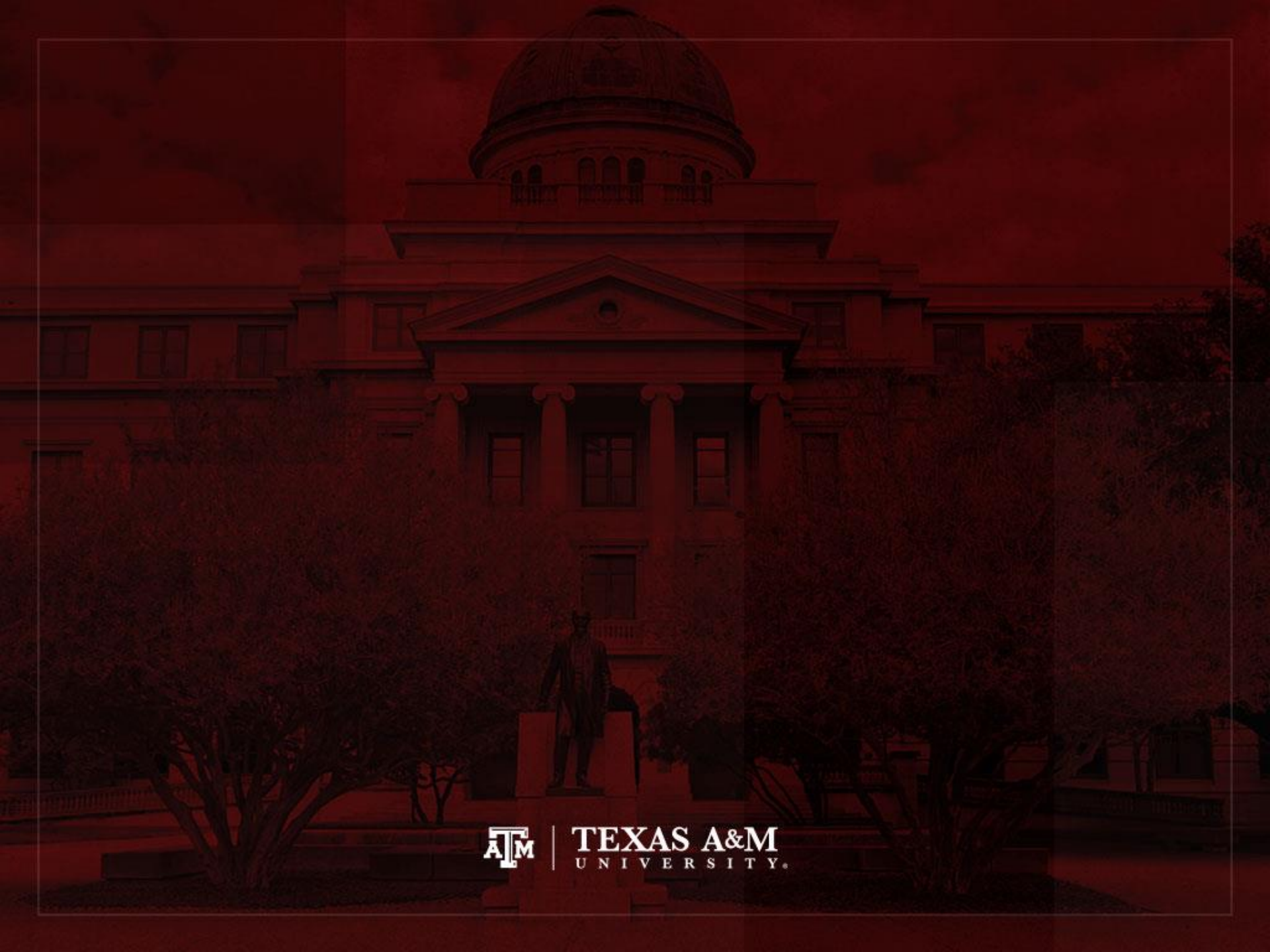
([https://gssdataexplorer.norc.org/gss\\_stdError](https://gssdataexplorer.norc.org/gss_stdError))

- VSTRAT and VPSU were created with a minimum of three respondents within a cell
  - If all cases are missing on a variable, you get an error message in Stata
  - “Missing standard error because of stratum with single sampling unit”
- It is recommended to utilize the “subpop” option for any subdomain analyses (e.g., for males)

```
svy, subpop(if sex==1): tab x
```

- You can also specify that strata with one sampling unit are “centered” at grand mean instead of stratum mean

```
svyset [weight=weightfinal], strata(vstrat) psu(vpsu) singleunit(centered)
```



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# Example: 2021 GSS in Stata (nominal-level variable)

`. tab sex`

respondents sex	Freq.	Percent	Cum.
male	1,736	44.06	44.06
female	2,204	55.94	100.00
Total	3,940	100.00	

`. tab sex [iweight=wtssnrps]`

respondents sex	Freq.	Percent	Cum.
male	1,904.2566	48.86	48.86
female	1,993.21543	51.14	100.00
Total	3,897.472	100.00	

`. svyset [weight=wtssnrps], strata(vstrat) psu(vpsu) singleunit(scaled)  
(sampling weights assumed)`

`. svy: tab sex`  
(running **tabulate** on estimation sample)

Number of strata = 9  
 Number of PSUs = 3,492  
 Number of obs = 3,940  
 Population size = 3,897.472  
 Design df = 3,483

responden ts sex	proportion
male	.4886
female	.5114
Total	1

Key: proportion = Cell proportion

# Example: 2021 GSS in Stata (ordinal-level variable)

. tab degree

r's highest degree	Freq.	Percent	Cum.
less than high school	246	6.14	6.14
high school	1,597	39.84	45.97
associate/junior college	370	9.23	55.20
bachelor's	1,036	25.84	81.04
graduate	760	18.96	100.00
Total	4,009	100.00	

. tab degree [iweight=wtssnrps]

r's highest degree	Freq.	Percent	Cum.
less than high school	480.972702	11.99	11.99
high school	1,891.6334	47.15	59.13
associate/junior college	452.656901	11.28	70.42
bachelor's	681.8664156	16.99	87.41
graduate	505.084448	12.59	100.00
Total	4,012.2139	100.00	

```
. svyset [weight=wtssnrps], strata(vstrat) psu(vpsu) singleunit(scaled)
(sampling weights assumed)
```

. svy: tab degree

(running tabulate on estimation sample)

```
Number of strata = 9
Number of PSUs = 3,543
Number of obs = 4,009
Population size = 4,012.2139
Design df = 3,534
```

r's highest degree	proportion
less than high school	.1199
high school	.4715
associate/junior college	.1128
bachelor's	.1699
graduate	.1259
Total	1

Key: proportion = Cell proportion

# Example: 2021 GSS in Stata (interval-ratio-level variable)

```
. sum conrinc
```

Variable	Obs	Mean	Std. dev.	Min	Max
conrinc	2,456	41722.79	39243.69	336	170912.6

```
. sum conrinc [iweight=wtssnrps]
```

Variable	Obs	Weight	Mean	Std. dev.	Min	Max
conrinc	2,456	2453.15509	37647.74	37376.88	336	170912.6

```
. svy: mean conrinc
```

(running mean on estimation sample)

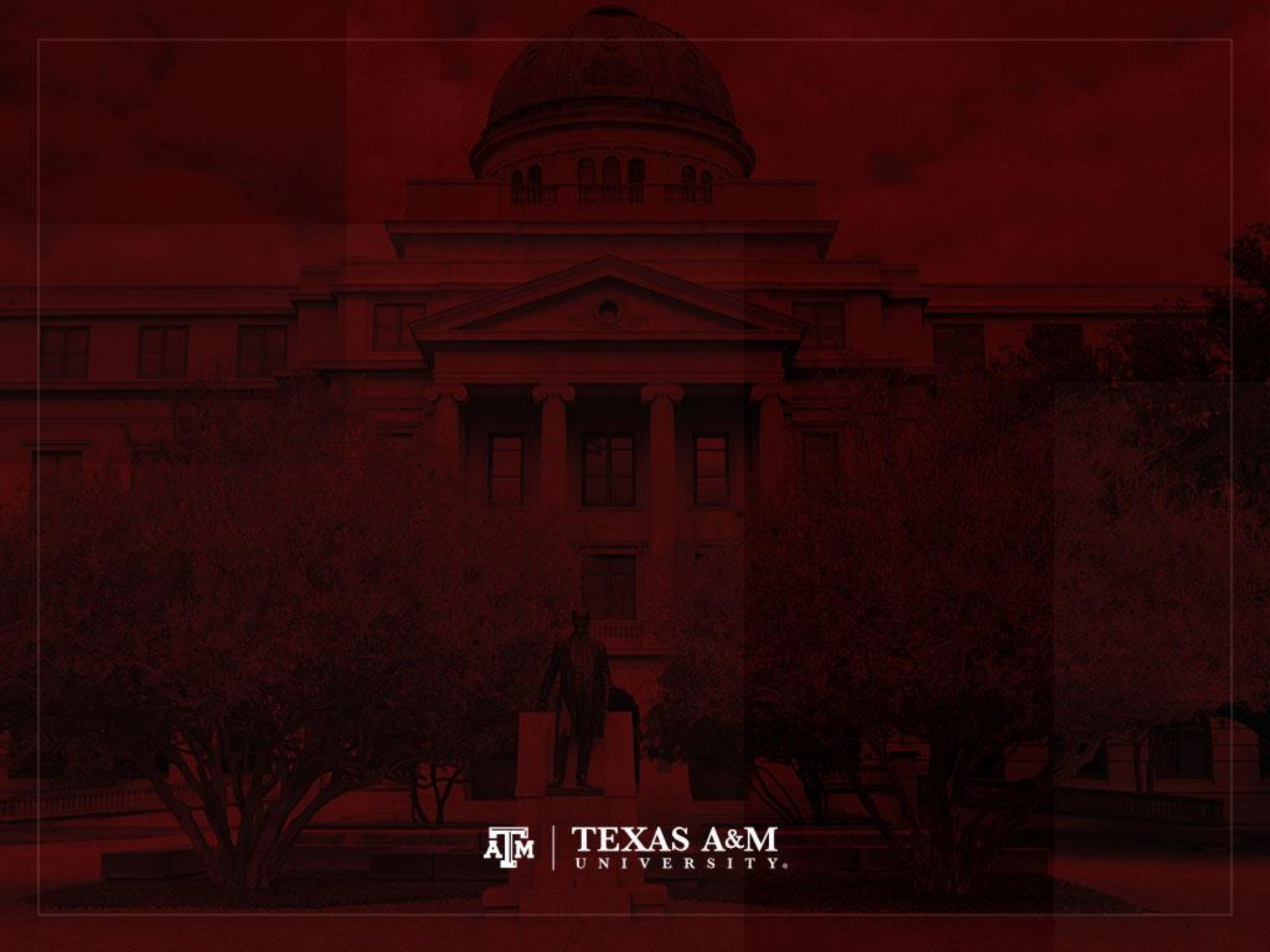
Survey: Mean estimation

```
Number of strata = 9          Number of obs = 2,456
Number of PSUs   = 2,241     Population size = 2,453.1551
Design df        =           Design df = 2,232
```

	Mean	Linearized std. err.	[95% conf. interval]	
conrinc	37647.74	850.3902	35980.1	39315.38

```
. estat sd
```

	Mean	Std. dev.
conrinc	37647.74	37376.87



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